

# An introduction to DYNARE

Michel Juillard (Banque de France and CEPREMAP)

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- 1 Introduction
- 2 A New Keynesian simple model
- 3 Writing a model in Dynare
- 4 Solving and simulating
- 5 Estimating a DSGE model with Dynare

- Microfoundations: agents optimize an objective function;
- Agents are forward looking: intertemporal optimization;
- Rational expectation hypothesis;
- General equilibrium: markets interact;
- Stochastic shocks push economy away from equilibrium;
- Endogenous dynamics bring the model back to equilibrium;
- We use numerical methods to approximate the solution of such models.

- 1 computes the solution of deterministic models (arbitrary accuracy),
- 2 computes first and second order approximation to solution of stochastic models,
- 3 estimates (maximum likelihood or Bayesian approach) parameters of DSGE models, for linear and non-linear models.
- 4 check for identification of estimated parameters
- 5 computes optimal policy
- 6 performs global sensitivity analysis of a model,
- 7 estimates BVAR and Markov-Switching Bayesian VAR models.

# A simple Neo-Keynesian model

- A simple model from [? ?]. Similar to [?] or [?].
- Main differences with RBC models:
  - ① imperfect competition
  - ② nominal rigidities
- Consequence: monetary policy matters

The representative household maximizes

$$\mathbb{E}_0 \sum_{s=0}^{\infty} \beta^s \left\{ \frac{(C_{t+s}/A_{t+s})^{1-\tau} - 1}{1-\tau} + \chi_M \ln \left( \frac{M_{t+s}}{P_{t+s}} \right) - \chi_H H_{t+s} \right\},$$

with:

$C_t$ : consumption index, in period  $t$

$A_t$ : level of technology

$\frac{M_t}{P_t}$ : real money balances,

$H_t$ : hours worked,

$\frac{1}{\tau}$ : elasticity of intertemporal substitution,

$\chi_M$ : scale factor

$\chi_H$ : scale factor.

# Household's budget constraint

$$P_t C_t + B_t + M_t - M_{t-} + T_t = P_t W_t H_t + R_{t-} B_{t-} + P_t D_t,$$

with

$P_t$ : price index for final goods,

$W_t$ : real wage,

$B_t$ : nominal government bond,

$T_t$ : lump-sum taxes,

$D_t$ : dividends received from monopolistic firms.

In addition, the usual transversality condition applies, excluding Ponzi schemes.

The representative firm producing the final good assembles a continuum of intermediate goods indexed by  $j \in [0, 1]$  in a perfectly competitive manner:

$$Y_t = \int_0^1 Y_t(j)^{-\nu} dj^{\frac{1}{1-\nu}},$$

with:

$Y_t$ : final good production index,

$Y_t(j)$ : good  $j$  production,

$\bar{\nu}$ : demand elasticity.



Under optimal behaviour, the price of the final good is

$$P_t = \int_0^{\infty} P_t(j)^{\frac{\nu-1}{\nu}} dj^{\frac{\nu}{\nu-1}},$$

with

- $P_t$ : price index of final good,
- $P_t(j)$ : price index of good  $j$ .

# Intermediate goods firms (I)

- Intermediate good  $j$  is produced by a monopolist with technology

$$Y(j) = A_t N_t(j),$$

where

$N_t(j)$ : labor input of firm  $j$ .

- Firms face a quadratic adjustment cost when they change their price:

$$AC(j) = \frac{\phi}{2} \left( \frac{P_t(j)}{P_{t-1}(j)} - \pi \right)^2$$

with

$\pi$ : steady state inflation rate

$\phi$ : adjustment cost parameter

## Intermediate goods firms (II)

Intermediate goods firms maximize the present value of future profits:

$$\mathbb{E}_t \sum_{s=0}^{\infty} \beta^s Q_{t+s} \frac{P_{t+s}(j)}{P_{t+s}} Y_{t+s}(j) - W_{t+s} N_{t+s}(j) - AC_{t+s}(j) \quad ,$$

with

$Q_{t+s}$ : marginal value of a unit of consumption good in period  $t + s$

The central bank stabilizes inflation and its policy is represented by an interest rate feedback rule:

$$R_t = R_{t-}^{\rho_R} \left( r \pi^* \left( \frac{\pi_t}{\pi^*} \right)^{\psi_1} \frac{Y_t}{Y_t^*}{}^{\psi_2} \right)^{-\rho_R} e^{\epsilon_{R,t}},$$

with

- $r$ : steady state real interest rate,
- $\pi_t$ : gross inflation rate ( $P_t/P_{t-}$ )
- $\pi^*$ : target inflation rate,
- $Y_t^*$ : potential output,
- $\epsilon_{R,t}$  monetary policy shock.

The government consumes a fraction  $\zeta_t$  of output:

$$G_t = \zeta_t Y_t,$$

with

$G_t$ : government expenditures.

The government levies a lump-sum tax  $T_t$  in order to balance the budget. The government budget constraint is

$$P_t G_t + R_{t-} B_{t-} = T_t + B_t + M_t - M_{t-} .$$

The following equilibrium relationships prevail:

$$Y_t = C_t + G_t + AC_t$$

$$H_t = N_t$$

There are three shocks in the model:

- 1 aggregate productivity  $A_t$
- 2 share of government expenditures in total output  $zeta_t$
- 3 monetary policy shock,  $\epsilon_{R,t}$

$\epsilon_{R,t}$  is serially uncorrelated.

$A_t$  evolves according to

$$\ln A_t = \ln \gamma + \ln A_{t-1} + \ln z_t,$$

and

$$\ln z_t = \rho_z \ln z_{t-1} + \epsilon_{z,t},$$

with

$\gamma$ : long run rate of growth of technology.



Define

$$g_t = \frac{1}{1 - \zeta_t}$$

and

$$\ln g_t = (1 - \rho_g) \ln g + \rho_g \ln g_{t-1} + \epsilon_{g,t},$$

with

$g$ : steady state share of government expenditures.

Potential output is defined as the output that would prevail in absence of nominal rigidities:

$$Y_t^* = (1 - \nu)^{\frac{1}{\tau}} A_t g_t.$$

# Dynamic recursive equilibrium (non-stationary)

$$1 = \beta \mathbb{E}_t \frac{C_{t+1}}{A_{t+1}} \frac{A_t}{C_t}^{-\tau} \frac{A_t}{A_{t+1}} \frac{R_t}{\pi_t},$$

$$1 = \frac{1}{\nu} \left[ 1 - \frac{C_t}{A_t} \right]^{\tau} + \phi (\pi_t - \pi) \left[ \left( 1 - \frac{1}{2\nu} \pi_t + \frac{\pi}{2\nu} \right) \right. \\ \left. - \phi \beta \mathbb{E}_t \frac{C_{t+1}}{A_{t+1}} \frac{A_t}{C_t}^{-\tau} \frac{Y_{t+1}}{A_{t+1}} \frac{Y_t}{A_t} (\pi_{t+1} - \pi) \pi_{t+1} \right],$$

$$Y_t^* = (1 - \nu)^{\frac{1}{\tau}} A_t g_t,$$

$$R_t = R_{t-1}^{\rho_R} r \pi^* \frac{\pi_t}{\pi^*} \psi_1 \frac{Y_t}{Y_t^*} \psi_2^{-\rho_R} e^{\epsilon_{R,t}},$$

$$\ln A_t = \ln \gamma + \ln A_{t-1} + z_t,$$

$$\ln z_t = \rho_z \ln z_{t-1} + \epsilon_{z,t},$$

$$\ln g_t = (1 - \rho_g) \ln g + \rho_g \ln g_{t-1} + \epsilon_{g,t},$$

# Detrending the model

- define detrended variables

$$c_t = \frac{C_t}{A_t}$$

$$y_t = \frac{Y_t}{A_t}$$

- steady state:

$$c = (1 - \nu)^{\frac{1}{\tau}}$$

$$y = g(1 - \nu)^{\frac{1}{\tau}}$$

$$r = \frac{\gamma}{\beta} R = r\pi^*$$

- define  $\hat{x}_t = \ln(\hat{x}_t/x)$  as the percentage deviation of variable  $x_t$ .

$$\begin{aligned}
 1 &= \mathbb{E}_t \left\{ e^{-\tau \hat{c}_{t+1} + \tau \hat{c}_t + \hat{R}_t - \hat{z}_{t+1} - \hat{\pi}_{t+1}} \right\}, \\
 \frac{1-\nu}{\nu \phi \pi^2} e^{\tau \hat{c}_t} - 1 &= e^{\hat{\pi}_t} - 1 \quad 1 - \frac{1}{2\nu} e^{\hat{\pi}_t} + \frac{1}{2\nu} \\
 &\quad - \beta \mathbb{E}_t \left\{ \left( e^{\hat{\pi}_{t+1}} - 1 \right) e^{-\tau \hat{c}_{t+1} + \tau \hat{c}_t + \hat{y}_{t+1} \hat{y}_t + \hat{\pi}_{t+1}} \right\} \\
 e^{\hat{c}_t - \hat{y}_t} &= e^{-g_t} - \frac{\phi \pi^2 g}{2} e^{\hat{\pi}_t} - 1^2 \\
 \hat{R}_t &= \rho_R \hat{R}_{t-} + (1 - \rho_R) (\psi_1 \hat{\pi}_t + \psi_2 (\hat{y}_t - \hat{g}_t)) + \epsilon_{R,t} \\
 \hat{z}_t &= \rho_z \hat{z}_{t-} + \epsilon_{z,t}, \\
 \hat{g}_t &= \rho_g \hat{g}_{t-} + \epsilon_{g,t},
 \end{aligned}$$

$$YGR_t = \gamma^{(Q)} + 100 (\hat{y}_t - \hat{y}_{t-1} + z_t)$$

$$INFL_t = \pi^{(A)} + 400 \hat{\pi}_t$$

$$INT_t = \pi^{(A)} + r^{(A)} + 4\gamma^{(Q)} + 400 \hat{R}_t$$

with

$\gamma^{(Q)}$ : rate of growth of output in percentage

$\pi^{(A)}$ : annual inflation rate in percentage

$r^{(A)}$ : annualized interest rate in percentage

Note that

$$\gamma = 1 + \frac{\gamma^{(Q)}}{100}$$

$$\beta = \frac{1}{1 + r^{(A)}/400}$$

$$\pi = 1 + \frac{\pi^{(A)}}{400}$$

```
var c R z pi y YGR INFL INT;  
varexo e_R e_g e_z;
```

```
parameters tau nu gbar rho_R rho_g rho_z  
             gamma_Q pi_A r_A kappa psi_1 psi_2;
```

```
tau = 2;  
nu = 0.1;  
kappa = 0.3;  
beta = 0.99;  
gbar = 1/0.85;  
rho_R = 0.5;  
rho_g = 0.8;  
rho_z = 0.66;  
pi_A = 4.0;  
r_A = 0.8;  
gamma_Q = 0.4;  
psi_1 = 1.5;  
psi_2 = 0.5;
```



```
model;  
# pi_bar = 1 + pi_A/400;  
# phi = tau*(1 - nu)/(nu*pi_bar^2*kappa);  
# gamma = 1 + gamma_Q/100;  
# beta = 1/(1+r_A/400);
```

## Dynare model file (IV)

```
exp(-tau*c(+1) + tau*c + R - z(+1) - pi(+1)) = 1;  
((1-nu)/(nu*phi*pibar^2))*(exp(tau*c) - 1) =  
  (exp(pi) - 1)*((1 - 1/(2*nu))*exp(pi) + 1/(2*nu))  
  - beta*(exp(pi(+1)) - 1)*exp(-tau*c(+1) + tau*c  
  + y(+1) - y + pi(+1));  
exp(c - y) = exp(-g)  
  - (phi*pibar^2*gbar/2)*(exp(pi) - 1)^2;  
R = rho_R*R(-1) + (1-rho_R)*(psi_1*pi + psi_2*(y - g))  
  + e_R;  
g = rho_g*g(-1) + e_g;  
z = rho_z*z(-1) + e_z;  
YGR = gamma_Q + 100*(y - y(-1) + z);  
INFL = pi_A + 400*pi;  
INT = pi_A + r_A + 4*gamma_Q + 400*R;  
end;
```

```
steady_state_model;  
c = 0;  
R = 0;  
z = 0;  
pi = 0;  
y = 0;  
g = 0;  
YGR = gamma_Q;  
INFL = pi_A;  
INT = pi_A + r_A + 4*gamma_Q;  
end;  
  
steady;
```

```
shocks;  
var e_R; stderr 0.003;  
var e_g; stderr 0.004;  
var e_z; stderr 0.004;  
end;  
  
stoch_simul(order=1);
```

## STEADY-STATE RESULTS :

c	0
R	0
z	0
pi	0
y	0
g	0
dgdp	0.4
pi_obs	4
R_obs	6.4

## MODEL SUMMARY

```
Number of variables:          9
Number of stochastic shocks:  3
Number of state variables:    4
Number of jumpers:           4
Number of static variables:   3
```

## MATRIX OF COVARIANCE OF EXOGENOUS SHOCKS

Variables	e_R	e_g	e_z
e_R	0.000009	0.000000	0.000000
e_g	0.000000	0.000016	0.000000
e_z	0.000000	0.000000	0.000016

## POLICY AND TRANSITION FUNCTIONS

	c	R
Constant	0	0
R(-1)	-0.282604	0.333974
g(-1)	0	0
z(-1)	0.294811	0.209595
y(-1)	0	0
e_R	-0.565209	0.667947
e_g	0	0
e_z	0.446684	0.317568

Approximated decision rule

$$\hat{c}_t = -0.283\hat{R}_{t-1} + 0.295\hat{z}_{t-1} - 0.565\epsilon_{R,t} + 0.447\epsilon_{z,t}$$

## THEORETICAL MOMENTS

VARIABLE	MEAN	STD. DEV.	VARIANCE
c	0.0000	0.0027	0.0000
R	0.0000	0.0031	0.0000
z	0.0000	0.0053	0.0000
pi	0.0000	0.0015	0.0000
y	0.0000	0.0072	0.0001
g	0.0000	0.0067	0.0000
dgdg	0.4000	0.7869	0.6192
pi_obs	4.0000	0.6120	0.3746
R_obs	6.4000	1.2366	1.5291



## VARIANCE DECOMPOSITION (in percent)

	e_R	e_g	e_z
c	43.91	0.00	56.09
R	47.29	0.00	52.71
z	0.00	0.00	100.00
pi	27.99	0.00	72.01
y	6.25	85.78	7.98
g	0.00	100.00	0.00
dgdp	6.96	28.71	64.33
pi_obs	27.99	0.00	72.01
R_obs	47.29	0.00	52.71

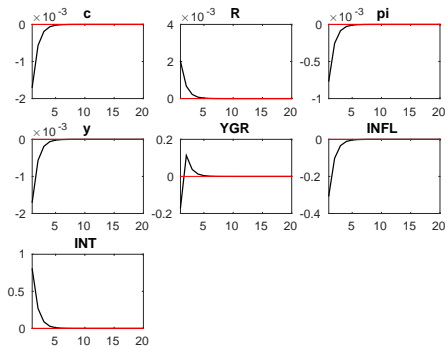
## MATRIX OF CORRELATIONS

Variables	c	R	z	pi
c	1.0000	0.0208	0.7270	0.9844
R	0.0208	1.0000	0.7016	0.1961
z	0.7270	0.7016	1.0000	0.8363
pi	0.9844	0.1961	0.8363	1.0000
y	0.3772	0.0078	0.2742	0.3713
g	0.0000	0.0000	0.0000	0.0000
dgdp	0.6942	0.3890	0.7619	0.7493
pi_obs	0.9844	0.1961	0.8363	1.0000
R_obs	0.0208	1.0000	0.7016	0.1961

## COEFFICIENTS OF AUTOCORRELATION

Order	1	2	3	4	5
c	0.4137	0.1907	0.0984	0.0558	0.0337
R	0.5872	0.3633	0.2316	0.1502	0.0982
z	0.6600	0.4356	0.2875	0.1897	0.1252
pi	0.4769	0.2536	0.1470	0.0902	0.0572
y	0.7450	0.5761	0.4532	0.3593	0.2859
g	0.8000	0.6400	0.5120	0.4096	0.3277
dgdp	0.1808	0.1482	0.1055	0.0707	0.0459
pi_obs	0.4769	0.2536	0.1470	0.0902	0.0572
R_obs	0.5872	0.3633	0.2316	0.1502	0.0982

# Impulse response function to monetary policy shock



# ESTIMATION

## Calibration:

- Models represents only some aspects of reality.
- Finding calibration that reproduces these aspects.
- Shortcoming: non indication of calibration uncertainty

## Bayesian estimation:

- Combining previous information (previous studies, microeconomic studies) with data.
- The posterior distribution displays uncertainty surrounding parameter estimates.
- Bayesian approach throw a bridge between calibration and classical methods.

- Experience shows that it is quite difficult to estimate a DSGE model by maximum likelihood.
  - 1 Data are not informative enough... The likelihood is flat in some directions (identification issue). This suggests that (when possible) we should use other sources of information.
  - 2 DSGE models are misspecified. When a DSGE is estimated by ML or with a “non informative” Bayesian approach (uniform priors) the estimated parameters are often found to be incredible. Using prior informations we can shrink the estimates towards sensible values.
- A related motivation is the relative lack of precision of ML. Prior information reduces the uncertainty.
- Finally the Bayesian approach allows easier comparison of (non nested) models.

- 1 priors specification as probability distributions,
- 2 computation of posterior distribution, on the basis of prior distribution and likelihood,
  - computing posterior mode,
  - simulating posterior distribution,
- 3 computing data marginal density useful for models comparison,
- 4 computing posterior predictive density
- 5 computing posterior distribution of IRFs, forecasts, etc . . .



- only some variables are observed,
- statistical model: unobserved components model,
- likelihood must be evaluated through a state space representation of the model and the Kalman filter,
- there must be at least as many shocks as observed variables,
- otherwise, the model suffers from stochastic singularity.

- Domain of definition:

Normal	$\mathbb{R}$
Uniform(a,b)	$[a, b]$
Gamma	$\mathbb{R}^+$
Beta	$[0, 1]$
Inverted gamma	$\mathbb{R}^+$
- mean prior: expected value,
- a small standard deviation means a tight prior,
- in Dynare, priors are orthogonal,
- implicit prior: unicity of stable equilibrium.

```
var c R z pi y g YGR INFL INT;  
varexo e_R e_g e_z;
```

```
parameters tau nu gbar rho_R rho_g rho_z gamma_Q pi_A  
             r_A kappa psi_1 psi_2;
```

```
nu =      0.1;  
beta =    0.99;  
gbar =    1/0.85;
```

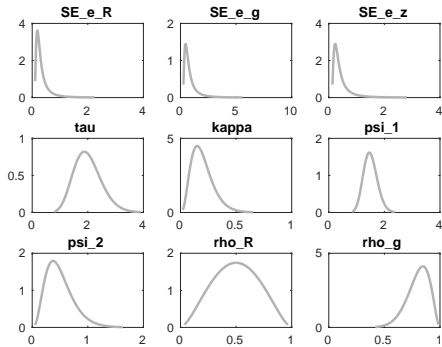
```
model;  
# pi_bar = 1 + pi_A/400;  
# phi = tau*(1 - nu)/(nu*pi_bar^2*kappa);  
# gamma = 1 + gamma_Q/100;  
# beta = 1/(1+r_A/400);
```

```
exp(-tau*c(+1) + tau*c + R - z(+1) - pi(+1)) = 1;  
((1-nu)/(nu*phi*pibar^2))*(exp(tau*c) - 1) =  
  (exp(pi) - 1)*((1 - 1/(2*nu))*exp(pi) + 1/(2*nu))  
  - beta*(exp(pi(+1)) - 1)*exp(-tau*c(+1) + tau*c  
  + y(+1) - y + pi(+1));  
exp(c - y) = exp(-g) - (phi*pibar^2*gbar/2  
                        *(exp(pi) - 1)^2;  
R = rho_R*R(-1) + (1-rho_R)*(psi_1*pi + psi_2*(y - g))  
  + e_R/100;  
g = rho_g*g(-1) + e_g/100;  
z = rho_z*z(-1) + e_z/100;  
YGR = gamma_Q + 100*(y - y(-1) + z);  
INFL = pi_A + 400*pi;  
INT = pi_A + r_A + 4*gamma_Q + 400*R;  
end;
```

```
steady_state_model;  
c = 0;  
R = 0;  
z = 0;  
pi = 0;  
y = 0;  
g = 0;  
YGR = gamma_Q;  
INFL = pi_A;  
INT = pi_A + r_A + 4*gamma_Q;  
end;
```

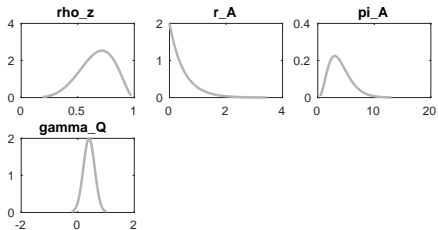
```
estimated_params;  
tau, gamma_pdf, 2.0, 0.5;  
kappa, gamma_pdf, 0.2, 0.1;  
psi_1, gamma_pdf, 1.5, 0.25;  
psi_2, gamma_pdf, 0.5, 0.25;  
rho_R, beta_pdf, 0.5, 0.2;  
rho_g, beta_pdf, 0.8, 0.1;  
rho_z, beta_pdf, 0.666, 0.15;  
r_A, gamma_pdf, 0.5, 0.5;  
pi_A, gamma_pdf, 4.0, 2.0;  
gamma_Q, normal_pdf, 0.4, 0.2;  
stderr e_R, inv_gamma_pdf, 0.4, inf;  
stderr e_g, inv_gamma_pdf, 1.0, inf;  
stderr e_z, inv_gamma_pdf, 0.5, inf;  
end;
```

# Priors plot (I)





# Priors plot (II)



```
varobs YGR INFL INT;  
  
// computing posterior mode  
estimation(datafile=as_data,first_obs=50,nobs=250,  
           mh_replic=0,mode_check);
```

# Dynare Output (I)

```
Initial value of the log posterior
                    (or likelihood): -812.3343
-----
f at the beginning of new iteration, 812.3343
Predicted improvement:      463.840722155
lambda =          1; f =      30459.8501185
lambda =    0.33333; f =      1092.5770840
lambda =    0.11111; f =       752.7353413
Norm of dx      0.30458
----
Improvement on iteration 1 =      59.598924800
...
Improvement on iteration 42 =      0.000000003
improvement < crit termination

Final value of minus the log posterior
                    (or likelihood):590.429395
```

## RESULTS FROM POSTERIOR ESTIMATION

parameters

	prior mean	mode	s.d.	prior	pstdev
tau	2.000	2.0854	0.3799	gamm	0.5000
kappa	0.200	0.1697	0.0253	gamm	0.1000
psi_1	1.500	1.4385	0.2320	gamm	0.2500
psi_2	0.500	0.3303	0.1226	gamm	0.2500
rho_R	0.500	0.5613	0.0422	beta	0.2000
rho_g	0.800	0.9529	0.0169	beta	0.1000
rho_z	0.666	0.5736	0.0290	beta	0.1500
r_A	0.500	0.6574	0.2722	gamm	0.5000
pi_A	4.000	3.9657	0.0338	gamm	2.0000
gamma_Q	0.400	0.3744	0.0762	norm	0.2000

standard deviation of shocks

	prior mean	mode	s.d.	prior	pstdev
e_R	0.400	0.1757	0.0093	invg	Inf
e_g	1.000	0.7684	0.0342	invg	Inf
e_z	0.500	0.5492	0.0451	invg	Inf

Log data density [Laplace approximation] is -621.416125.

# Simulating the posterior distribution

- The posterior distribution doesn't have an analytical representation.
- It is possible to simulate a sample of parameter values, representative of the posterior distribution.
- Metropolis is a MCMC algorithm with acceptance/rejection of proposed parameter values randomly drawn.
- The average acceptance must be monitored. A target of about 25% is considered optimal.
- The average acceptance ratio can be tuned with the **mh\_jscale** parameter of the **estimation** command.

```
estimation(datafile=as_data,first_obs=50,nobs=250,  
           mode_compute=0,mode_file=as_est_mode,  
           mh_replic=100000,mh_nblocks=1,mh_jscale=0.7,  
           fast_kalman_filter,  
           bayesian_irf,forecast=12);
```

- Only some variables are observed (here YGR INFL INT).
- The Kalman smoother finds best values for other variables and for shocks.
- In absence of measurement errors, the smoothed value of the observed variables is equal to the observation.



```
Estimation::mcmc: Current acceptance ratio per chain:  
Chain 1: 21.08%
```

```
ESTIMATION RESULTS
```

```
Log data density is -621.301701.
```

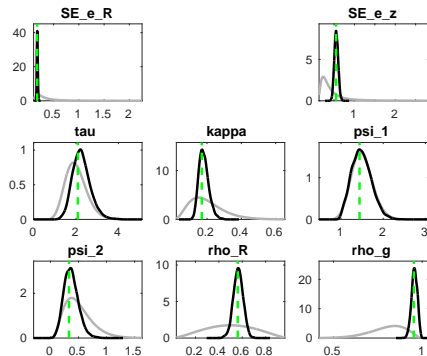
## parameters

	prior mean	post. mean	90% HPD interval	
tau	2.000	2.2520	1.5927	2.9147
kappa	0.200	0.1815	0.1363	0.2274
psi_1	1.500	1.4870	1.0970	1.8468
psi_2	0.500	0.3673	0.1619	0.5713
rho_R	0.500	0.5651	0.4986	0.6323
rho_g	0.800	0.9559	0.9295	0.9824
rho_z	0.666	0.5773	0.5297	0.6250
r_A	0.500	0.6813	0.2418	1.1119
pi_A	4.000	3.9629	3.9110	4.0173
gamma_Q	0.400	0.3676	0.2473	0.4929

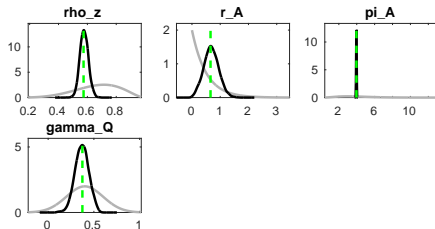
standard deviation of shocks

	prior mean	post. mean	90% HPD interval	
e_R	0.400	0.1782	0.1630	0.1940
e_g	1.000	0.7777	0.7208	0.8352
e_z	0.500	0.5530	0.4755	0.6275

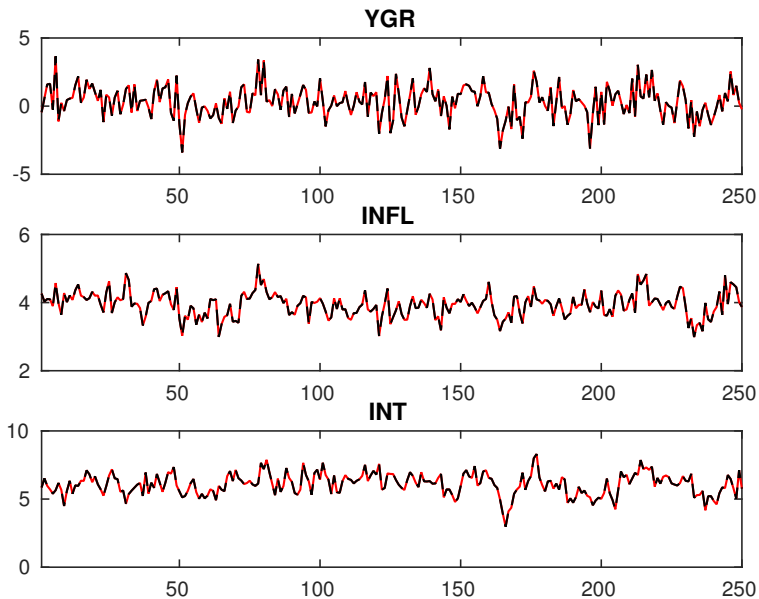
# Prior-Posterior-Mode (I)



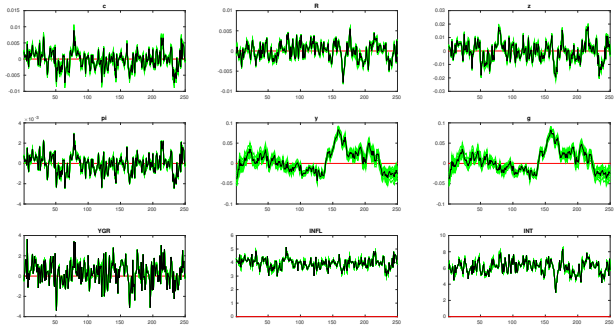
# Prior-Posterior-Mode (II)



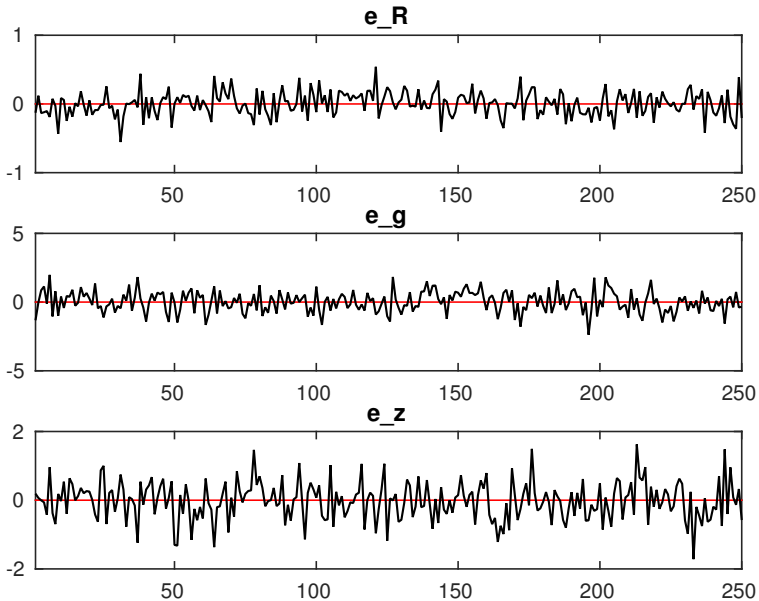
# Smoothed observed variables



# Smoothed variables

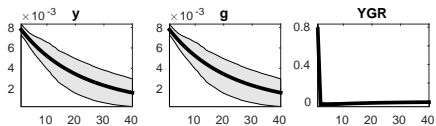


# Smoothed shocks

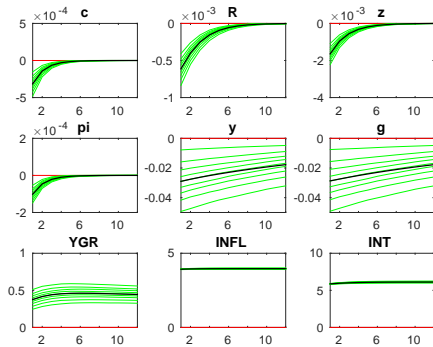




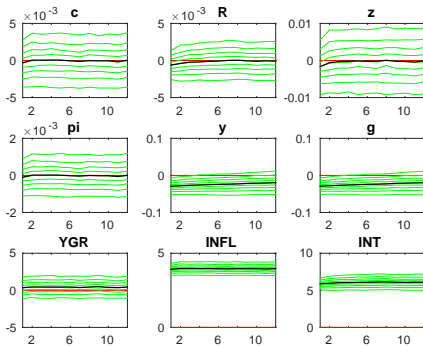
# Posterior distribution of IRFs



# Posterior distribution of forecast mean

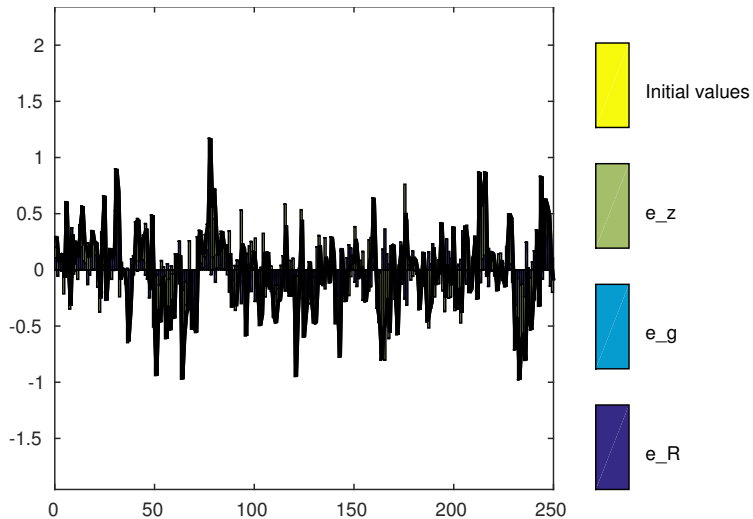


# Posterior distribution of forecast point



- Because the reduced form of the model is a state space, it is possible to decompose the dynamics of each endogenous variables as the sum of initial conditions and the cumulative effects of the shocks
- In real applications, this decomposition can be compared to a narrative of events.

# Historical shock decomposition of inflation



- Numerical optimization for finding the mode not always works:

```
POSTERIOR KERNEL OPTIMIZATION PROBLEM!
```

```
(minus) the hessian matrix at the "mode" is not positive definite!
```

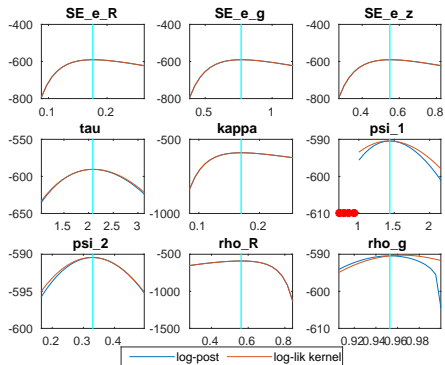
```
=> posterior variance of the estimated parameters are not positive.
```

```
You should try to change the initial values of the parameters using  
the estimated_params_init block, or use another optimization routine.
```

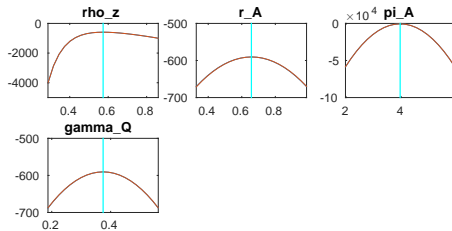
```
Warning: The results below are most likely wrong!
```

- Check for parameter standard deviation of 0
- If needed use option `mode_check`

# Mode check



# Mode check



— log-post — log-lik kernel



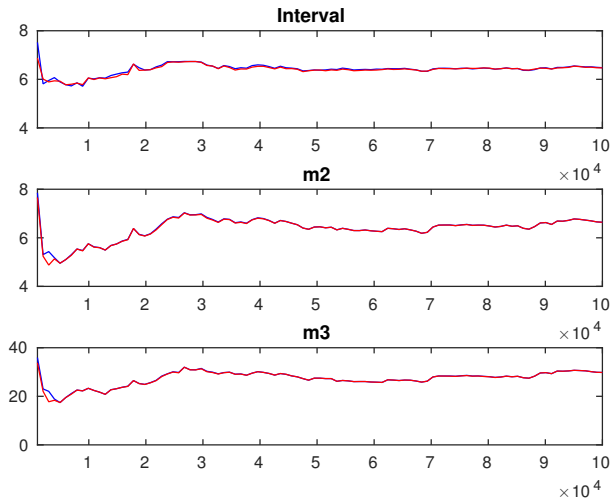
- Numerical optimization for finding the mode not always works:

```
POSTERIOR KERNEL OPTIMIZATION PROBLEM!
```

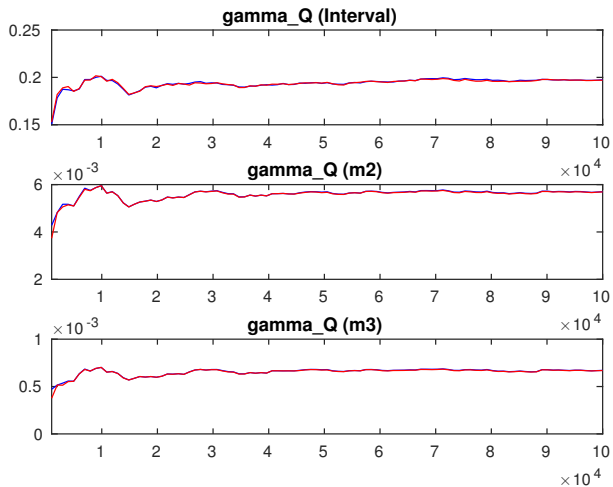
```
(minus) the hessian matrix at the "mode" is not positive definite!  
=> posterior variance of the estimated parameters are not positive.  
You should try to change the initial values of the parameters using  
the estimated_params_init block, or use another optimization routine.  
Warning: The results below are most likely wrong!
```

- Check for parameter standard deviation of 0
- If needed use option `mode_check`
- Assess convergence of Metropolis sampling: Brooks and Gelman (1998) convergence tests

# Multivariate convergence tests



# Univariate convergence tests



- Data must be consistent with the variables of the model (stationarity units)
- Check that a model simulated under the prior produces moments roughly comparable with the empirical moments of the data
- Start with a small model and add mechanisms one by one. This makes it easier to find mistakes. You can use better initial values for the parameters.